SUPPLEMENT TO "THE VALUE OF FREE WATER: ANALYZING SOUTH AFRICA'S FREE BASIC WATER POLICY" (*Econometrica*, Vol. 83, No. 5, September 2015, 1913–1961)

BY ANDREA SZABÓ

This supplement, not intended for publication, contains further details on the data and analysis presented in the paper.

S1. DETAILS ON THE SURVEY

THE GOAL OF THE SURVEY was to collect information on water usage and household demographics to complement the administrative data provided by Odi Water. The survey included 21 questions about, for example, the number and type of water-using fixtures used by each household, the lot size and landscape characteristics, relevant irrigation methods and habits, other exterior uses of water, water conservation actions taken, and household demographics such as family size, education levels, and income.

The survey was entirely funded by the University of Houston. It was approved by the Human Subject Committee of the University of Houston, and was conducted in accordance with the standards of that institution regarding the ethical treatment of human subjects. Participation in the survey was voluntary and respondents could stop participating in the survey at any time. Each questionnaire took about 30 minutes to complete. Only adults between the ages of 18 and 65 were asked to participate.

S1.1. Sampling

The objective of the sampling design was to yield a sample that is representative of the surveyed population, the residential consumers of Odi Water, based on information that was available prior to the survey. This included monthly water consumption, indigent status, whether the consumer was restricted, and the supply area. These variables were used to define 96 strata in the population, and a stratified random sample was taken. At the time of the survey, December 2008 was the last monthly billing cycle for which information on Odi Water's consumers was available, and the sampling was therefore designed to be representative of that population. To reach my target sample size of 1000, I took an initial sample of 1000 households. The survey team was instructed to interview each of the 1000 households in this sample. If a household could not be reached after two attempts or if the respondent declined to answer, it was replaced by a random household from the un-sampled population of consumers belonging to the same stratum as the original household.

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If that household could not be reached or declined to answer, it would be replaced by a third household from the same stratum, and so on. This procedure ensured that the resulting sample was representative of the surveyed population. In practice, 177 households were replaced once, 29 twice, and 9 three times before the data collection ended. The final sample size is 1000 households.

S1.2. Data Collection

The survey was carried out by Impact Research International, Inc. (IRI), a survey and market research company based in Pretoria, under the supervision of the author. Surveyors with extensive local experience were trained by IRI and the author specifically for the purpose of this project. These surveyors visited each household at their home address and asked and recorded the questions to the questionnaire provided by the author. Interviews were performed in either English or a local language (all interviewers were fluent in all the relevant local languages), and the questionnaire itself was available in all the relevant languages. The completed questionnaires went through a quality check by the staff at IRI and by the author, and were entered into a computer database. After a pilot survey on November 25, 2010, data collection took place between November 29, 2010 and February 2, 2011.

Throughout the process, we ensured that the survey was performed independently from the personnel at Odi Water or the municipality. Surveyors introduced themselves to respondents as working under contract for the University of Houston. The separation of individual consumption information (including account numbers) and survey responses was maintained throughout the survey. The final database identifies respondents by a code generated by the author, and does not contain the billing account number, service address, or name of the customers.

S1.3. Sample Properties

Table S-I contains the main characteristics of the sample based on the latest information in Odi Water's records (December 2008). Approximately 11% of the respondents have indigent status and receive a free water allowance of 12 kl per month. Nineteen percent of the households are restricted. Average household consumption is 12.95 kl per month. The respondents live in three distinct supply areas based on Odi Water's definition. The table also shows the composition of surveyed population. As can be seen, the sample is representative with respect to the household characteristics available prior to the survey.

TABLE S-I

| | All Consumers December 2008 | Survey Full Sample |
|--------------------------------|--------------------------------|-----------------------|
| Indigent (%) | 11 | 11 |
| Restricted (%) | 19 | 19 |
| Sanitation (%) | 85 | 86 |
| Supply area 1 (%) | 31 | 30 |
| Supply area 2 (%) | 21 | 21 |
| Supply area 3 (%) | 48 | 49 |
| Average consumption (kl/month) | 12.95 | 12.1 |
| Monthly consumption | | |
| 4 kl or less (%) | 25.6 | 25.2 |
| 5–8 kl (%) | 24.7 | 25.3 |
| 9–14 kl (%) | 23.9 | 26.5 |
| 15 or more kl (%) | 25.8 | 23 |
| Number of observations | 46,214 | 1000 |

SAMPLE PROPERTIES^a

 $^{\rm a}{\rm The}$ table compares the distribution of consumer characteristics in the population and the survey sample.

S2. PHOTOS FROM THE PROJECT

Figures S2 and S2 illustrate the study area. Typical properties in the area are relatively small; the average household size is 4.3 persons. Many housing units are uniform single-family buildings provided by the government with limited modifications made by the residents. Thus, living conditions are fairly similar within the sample.



FIGURE S1.—Water meter.



FIGURE S2.—Typical housing unit and infrastructure in the sample.

Each participating household has water-using sanitation and tap(s) inside the house or outside on the property. Consumption is metered individually by a meter located on the property and easily accessible to the household.

S3. ADMINISTRATIVE DATA

S3.1. Water Consumption Data

This section describes cleaning of the administrative consumption data and how the final data set was generated. The final data set includes 3,036,871 monthly observations, and was generated as follows. I dropped from the data Odi Water's commercial and institutional consumers (5.6% of the data). I also dropped accounts showing zero consumption.¹ The employees of Odi Water inspect the water meter each month at meter-reading. If there is a problem with the meter, employees record the code of the problem which is also included in my data set.² In addition, the meter reader tests the tap for any water leaks and reports the problem. I drop observations with *any* problems reported (1.7%). Because of the regular quality checks, illegal tapping in this area is virtually non-existent, in contrast to many other developing countries or even other parts of South Africa.

Based on my conversations with Odi Water officials, the utility has difficulty distinguishing commercial and residential consumers if the consumer is running a small business from home. These small businesses include hair salons, car washing facilities, small restaurants, etc. Since Odi Water made efforts to identify these households and re-categorize them as commercial units, there are several account numbers whose status changed from domestic to commercial during the observed period. I drop the entire accounts with changing status from the sample (less than 0.2%). I also drop areas which are categorized as agricultural since water pricing and supply is different than in the rest of the supply areas (3.7%). Lastly, I drop observations where monthly consumption is higher than 50 kl, which is four times the average consumption (an additional 3.4%). These consumption levels are most likely associated either with unreported leaks or with commercial activities not yet categorized as such by Odi Water. In the final data set, average monthly consumption is 13.2 kl.

Table S-II presents detailed information on the distribution of consumption across all tariff schedules in the data. Twenty-eight point three percent of the

| Tariff Segment | % of Consumers | Ν |
|----------------------------|----------------|---------|
| Segment 1 (1–6 kl) | 28.32 | 859,931 |
| Segment 2 (7–12 kl) | 29.26 | 888,575 |
| Setment 3 (13–18 kl) | 18.86 | 572,784 |
| Segment 4 (19–24) | 10.73 | 325,789 |
| Segment 5 (25–30) | 5.88 | 178,524 |
| Segment 6 (31–42) | 5.08 | 154,219 |
| Other segments $(43 + kl)$ | 1.88 | 57,049 |

| TABLE S-II |
|--|
| DISTRIBUTION OF CONSUMERS BY TARIFF SEGMENT ^a |

^aThe table shows the distribution of consumption levels in the population across eight tariff years. A tariff segment is a consumption interval that is assigned the same unit price in the tariff schedule. N = 3,036,871.

¹In some cases, zero consumption refers to vacant land, closed accounts, etc., a total of 489,959 monthly observations.

²These problems include the following: dirty dial, meter covered, meter stuck, meter damaged, meter dial is missing, meter tampered with, meter obstructed, water leak, or meter removed.

| Water Consumption in kl | Sanitation Multiplier |
|-------------------------|-----------------------|
| 0–6 | 0.98 |
| 7–12 | 0.9 |
| 13–18 | 0.75 |
| 19–24 | 0.6 |
| 25-30 | 0.52 |
| 31-42 | 0.1 |
| 42+ | 0.01 |

TABLE S-IIISanitation Multiplier, 2002–2009a

^aThe table shows sanitation multipliers for different price segments. These determine the quantity after which sanitation charges are billed.

households consume below 6 kiloliters, which is the free allowance under most tariff schedules. Seventy-six point five percent of all observations are concentrated on the first three price blocks (up to 18 kl).

S3.2. Tariffs

Table S-III shows the sanitation multipliers included in the prices. Table S-IV presents summary statistics of the tariff schedules.

| Variable | Units | Mean | Std. Dev. | Min | Max |
|---------------------------|---------|-------|-----------|------|-------|
| Marginal price in block 1 | Rand/kl | 1.59 | 2.89 | 0 | 7.18 |
| Marginal price in block 2 | Rand/kl | 9.96 | 1.93 | 3.67 | 11.78 |
| Marginal price in block 3 | Rand/kl | 10.54 | 1.34 | 6.47 | 11.86 |
| Marginal price in block 4 | Rand/kl | 10.38 | 1.37 | 6.53 | 12.28 |
| Marginal price in block 5 | Rand/kl | 10.39 | 1.53 | 6.61 | 12.99 |
| Marginal price in block 6 | Rand/kl | 9.10 | 1.72 | 6.69 | 12.02 |
| Marginal price in block 7 | Rand/kl | 9.17 | 1.97 | 6.76 | 12.86 |
| Marginal price in block 8 | Rand/kl | 11.32 | 1.46 | 8.96 | 12.91 |
| Water quantity at kink 1 | kl | 6.00 | 0.05 | 6 | 12 |
| Water quantity at kink 2 | kl | 12.00 | 0.05 | 12 | 18 |
| Water quantity at kink 3 | kl | 18.00 | 0.05 | 18 | 24 |
| Water quantity at kink 4 | kl | 24.83 | 3.77 | 24 | 42 |
| Water quantity at kink 5 | kl | 31.93 | 8.80 | 30 | 72 |
| Water quantity at kink 6 | kl | 42.00 | 0.24 | 42 | 72 |
| Water quantity at kink 7 | kl | 75.68 | 7.26 | 72 | 90 |

 TABLE S-IV

 Summary Statistics, Tariff Structure, 2002–2009, 2008 Rand^a

^aThe table shows summary statistics of the price schedules (marginal prices and kink points) observed throughout the study period.

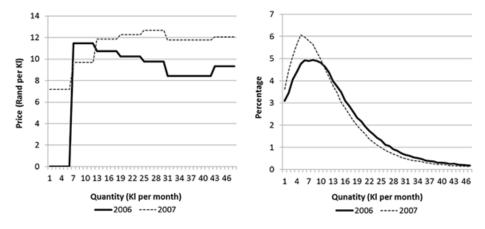


FIGURE S3.—Policy change for non-indigent households, 2006–2007. *Notes*: The left panel shows the change in the marginal prices for non-indigent households. The right panel shows the corresponding change in the consumption distribution.

Figures S3 and S4 show the policy change in 2007, when the utility created separate tariff structures for registered indigent households. The left panels show the change in the tariffs. The right panels show the corresponding change in the consumption distribution.

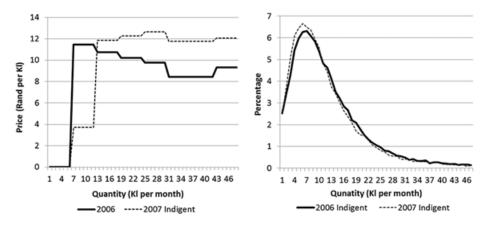


FIGURE S4.—Policy change for indigent households, 2006–2007. *Notes*: The left panel shows the change in the marginal prices for indigent households. The right panel shows the corresponding change in the consumption distribution.

S4. PATTERNS IN THE DATA

S4.1. Distribution of Consumption Over Time

Figures S5–S7 present various comparisons of consumption distributions over time. The figures compare the distribution of quarterly average consumption with the same quarter of the previous tariff year. Tariffs are revised starting in July of every year, so quarter 1 refers to July–September. Comparing the same quarters across years helps filter out any seasonality of water consumption in the raw data. Figure S5 focuses on tariff years 6 and 7, when the biggest tariff change occurred. Figure S8 shows the monthly distribution for these two years. These figures provide suggestive evidence that consumers respond to revisions of the price schedule in successive tariff years, perhaps with a lag of a month or two. Figures S9 and S10 show the distribution across price blocks by indigent status.

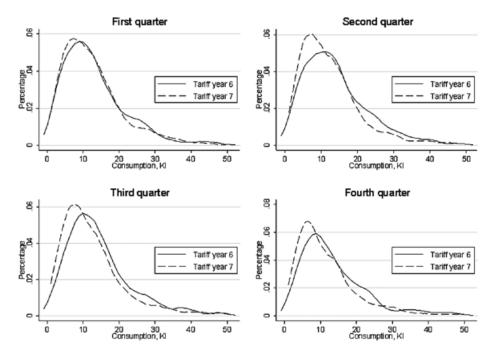


FIGURE S5.—Distribution of quarterly average consumption before and after the 2007 policy change. *Notes*: The figure presents densities of average consumption levels for each quarter of tariff years 6 and 7. Quarter 1: July–September, Quarter 2: October–December, Quarter 3: January–March, Quarter 4: April–June.

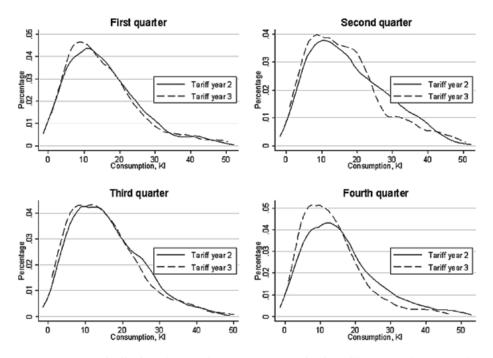


FIGURE S6.—Distribution of quarterly average consumption in tariff year 2 and 3. *Notes*: The figure presents densities of average consumption levels for each quarter of tariff years 2 and 3. Quarter 1: July–September, Quarter 2: October–December, Quarter 3: January–March, Quarter 4: April–June.

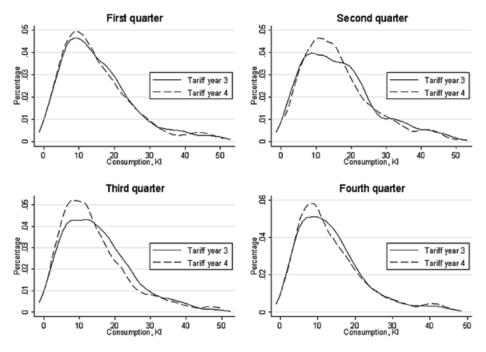


FIGURE S7.—Distribution of quarterly average consumption in tariff year 3 and 4. *Notes*: The figure presents densities of average consumption levels for each quarter of tariff years 3 and 4. Quarter 1: July–September, Quarter 2: October–December, Quarter 3: January–March, Quarter 4: April–June.

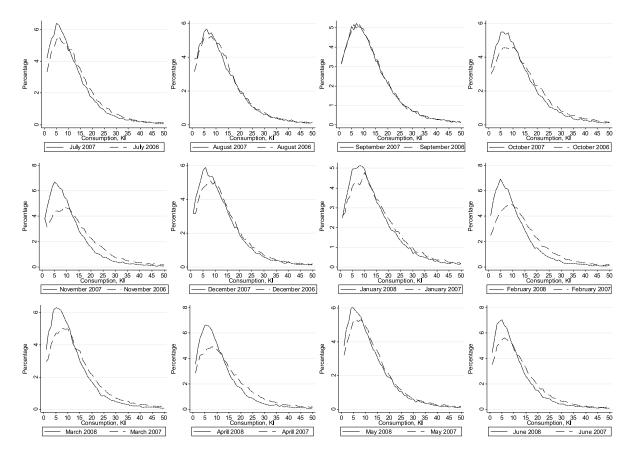


FIGURE S8.—Monthly consumption distribution in tariff year 6 and 7. *Note*: The figure presents densities of consumption levels in each month in tariff year 6 and 7.

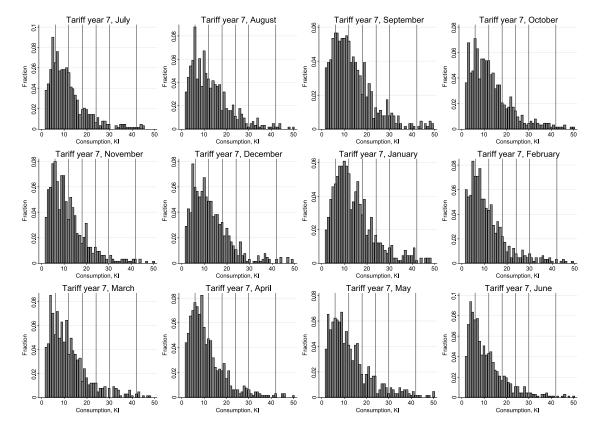


FIGURE S9.—Consumption distribution across price blocks, tariff year 7, only non-indigent households. *Notes*: The figure presents histograms of consumption levels in each month of tariff year 7. Each bar corresponds to a 1 kl increment in consumed quantity. Vertical lines represent the kink points of each tariff schedule.

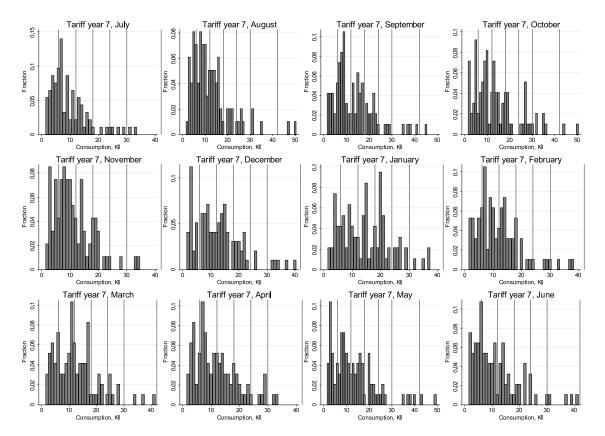


FIGURE S10.—Consumption distribution across price blocks, tariff year 7, only indigent households. *Notes*: The figure presents histograms of consumption levels in each month in tariff year 7. Each bar corresponds to a 1 kl increment in consumed quantity. Vertical lines represent the kink points of each tariff schedule.

S4.2. Detailed Regression Table

Column (1) of Table S-V shows the OLS demand estimate referred to in Section 3.2 in the main text. Columns (2)–(5) present the detailed regression output corresponding to the first row of Table II in the main text.

| Variables | (1) | (2) | (3) | (4) | (5) |
|------------------------------------|-------------------|-------------------|-------------------|-------------------|-------------------|
| Price | 1.948 (0.009) | -0.620 (0.024) | -0.608 (0.024) | -0.551 (0.023) | -0.554 (0.023) |
| Income $\times 10^{-4}$ | 1.151 (0.075) | 2.307 (0.112) | 1.081 (0.150) | 0.230 (0.146) | 0.875 (0.118) |
| Indigent | 2.531 (0.075) | -2.172 (0.143) | -2.396 (0.141) | -1.503 (0.137) | -1.262 (0.134) |
| Sanitation | -3.337 (0.119) | -0.366 (0.176) | -0.012 (0.177) | -1.095 (0.175) | 2.543 (0.157) |
| Average max daily temperature (°F) | 0.103 (0.005) | 0.105 (0.007) | 0.105 (0.007) | 0.104 (0.007) | 0.104 (0.007) |
| Supply area 1 | -1.412 (0.071) | -1.965 (0.100) | -1.749 (0.099) | -1.449 (0.098) | |
| Supply area 2 | -2.065 (0.099) | -6.284 (0.142) | -5.982 (0.141) | -4.778 (0.144) | |
| Restricted | 1.128 (0.079) | 1.058 (0.118) | 0.882 (0.116) | 1.167 (0.113) | 1.369 (0.114) |
| Number of flush toilets | | | | 0.633 (0.095) | |
| Number of standpipes | | | | -0.474 (0.046) | |
| Number of bathtubs | | | | 1.582 (0.093) | |
| Number of showers | | | | 0.684 (0.141) | |
| Number of kitchen taps | | | | -0.427 (0.093) | |
| Number of bathroom taps | | | | 1.381 (0.069) | |
| Washing machine | | | | 0.680 (0.093) | 0.781 (0.091) |
| Lawn area | | | | 1.081 (0.104) | |

TABLE S-V WATER DEMAND REGRESSIONS^a

(Continues)

| Variables | (1) | (2) | (3) | (4) | (5) |
|---|---------|---------|-------------------|-------------------|------------------|
| Flower garden | | | | -0.554 | |
| | | | | (0.100) | |
| Vegetable garden | | | | 1.107 | |
| Winton impostion | | | | (0.118) 0.336 | |
| Winter irrigation | | | | (0.058) | |
| Summer irrigation | | | | -0.226 | |
| | | | | (0.050) | |
| Washes car | | | | -0.085 | |
| | | | | (0.050) | |
| Some high school | | | 1.635 | 1.948 | |
| | | | (0.172) | (0.167) | |
| High school graduate | | | 1.029 (0.162) | 1.043 (0.156) | |
| Some higher education | | | 1.288 | 0.918 | |
| Some night education | | | (0.182) | (0.176) | |
| Completed higher education | | | 3.138 | 2.054 | |
| 1 0 | | | (0.206) | (0.196) | |
| Number of adults | | | 1.076 | 0.983 | |
| | | | (0.036) | (0.036) | |
| Number of teens | | | -0.114 | 0.109 | |
| Number of children | | | (0.047) | (0.048) | |
| Number of children | | | -0.041 (0.046) | -0.104 (0.046) | |
| Number of people working outside the home | | | -0.077 | -0.478 | |
| | | | (0.067) | (0.066) | |
| Number of persons on the property | | | | | 0.391 |
| | | | | | (0.020) |
| Outdoor water usage | | | | | 0.438 |
| | | | | | (0.043) |
| Bathtub or shower | | | | | 4.102 (0.094) |
| Completed high school or higher | | | | | -0.431 |
| completed light school of light | | | | | (0.095) |
| Constant | -1.102 | 9.857 | 5.779 | 4.156 | 0.887 |
| | (0.357) | (0.531) | (0.547) | (0.533) | (0.506) |

TABLE S-V—Continued

^aThe dependent variable is monthly consumption in kl. 'Price' is the average price corresponding to this consumption. Column (1) is an OLS regression, (2)–(5) are IV regressions where price is instrumented by the marginal prices at the six most common kink points (6, 12, 18, 24, 30, and 42 kl). 'Supply area' denotes a geographic area. 'Outdoor water usage' indicates how many of the following the household owns (0–3): vegetable garden, flower garden, lawn area. 'Bathtub or shower' is 1 if the household owns a bathtub, a shower, or both, and 0 otherwise. Robust standard error in parentheses. N = 63,178.

S4.3. Regressions Using Lagged Tariffs

This section presents further evidence on consumers' responsiveness to price changes. Specifically, for each month, I compute the average price faced by the household based both on the current tariff schedule as well as the tariff schedule from the previous year. I then regress consumption on the current and the previous average price, using the marginal prices of each tariff schedule as instruments, and controlling for household characteristics. I do this for all months, as well as taking averages within each quarter and restricting the sample to the same quarter in every tariff year. Both current and past average prices are instrumented with the marginal prices (as described in the main text, Section 3.2).

Table S-VI shows the coefficient estimates on current and lagged average prices from these regressions (in each case, the sample contains only nonindigent households with sanitation). When all months are included, the current and lagged average prices are both significant with the expected negative sign. The coefficient on the current price is twice as large as the one on the lagged price, indicating that consumption responds more to the current tariff schedule. A similar pattern is observed when restricting the sample to the first quarter in every tariff year. In the second quarter, the coefficient on the lagged price is only 1/5th of the current price coefficient, and it is not statistically significant. The results in the fourth quarter show a similar picture, with the lagged price coefficient close to zero. The third quarter is the only one showing a larger effect of the lagged price, but the coefficient on the current price becomes positive in this case, making this particular regression hard to interpret. Overall, these patterns echo those seen in the raw data, with consumption becoming more responsive to the new tariff later in the tariff year.

| | All | Q1 | Q2 | Q3 | Q4 |
|-------------------|----------------|----------------|---------------|--------------|----------------|
| Price in <i>t</i> | -0.398^{***} | -0.329^{***} | -0.419^{**} | 0.041 | -0.314^{***} |
| | (0.029) | (0.092) | (0.196) | (0.132) | (0.087) |
| Price in $t - 1$ | -0.332^{***} | -0.202^{*} | -1.101^{**} | -0.458^{*} | -0.275 |
| | (0.043) | (0.112) | (0.442) | (0.252) | (0.283) |

TABLE S-VI LAGGED TARIFF REGRESSIONS^a

^aThe dependent variable is monthly consumption in kl. 'Price' is the average price corresponding to this consumption based on the current (*t*) or previous (t - 1) price schedule. Prices are instrumented by the marginal prices at the six most common kink points (6, 12, 18, 24, 30, and 42 kl). Q1–Q4 restrict the sample to observations from specific quarters. The estimation sample contains non-indigent households with sanitation. Robust standard errors in parentheses. N = 63,178.

S5. DETAILS ON THE ESTIMATION PROCEDURE

S5.1. Steps of the Estimation Procedure

The demand estimation procedure described in the paper is computationally complex. The following describes the step by step instructions to estimate the demand function in the case of a mixture of increasing and decreasing tariffs. The procedure can be implemented in MATLAB or using similar software.

The following steps should be iterated from initial starting values for the parameters $(\alpha, \gamma, \delta, \sigma_{\eta}, \sigma_{\varepsilon})$ using any minimization procedure until convergence is achieved.

1. Compute the following for each individual:

(a) For each tariff segment k: Y_k^0 , $w_k^0 = Z\delta + \alpha P_k + \gamma Y_k^0$, and $\theta_{jk} = \bar{w}_j - w_k^0$, as well as $\bar{\eta} \equiv \min_k (-w_k^0 - \frac{\alpha}{\gamma})$.

(b) Compare segments to segments. For each pair of segments *i* and *j*: η_{ij}^s which solves $V(P_i, Y_i^0) = V(P_j, Y_j^0)$. Given the functional forms, η_{ij}^s has a closed form solution and is unique.

(c) Compare convex kink points to convex kink points. For each pair of convex kink points *i* and *j*: η_{ij}^k which solves $U(\bar{w}_i) = U(\bar{w}_j)$. Given the functional forms, η_{ij}^k has a closed form solution and is unique.

(d) Compare segments to convex kink points. For each convex kink *i* and segment $j \neq i \pm 1$: $u_{ij}^L < u_{ij}^U$ which are the two roots of the equation $U(\bar{w}_i) = V(P_j, Y_j^0)$ in η . Given the functional forms, for given *i* and *j*, there are a maximum of two roots which can be computed numerically (no closed form solution).

To make sure that a root exists, I first compute the value of η for which $\partial V/\partial \eta = \partial U/\partial \eta$. Equation $U(\bar{w}_i) = V(P_j, Y_j^0)$ has two roots iff at this value of $\eta, V_j < U_i$, in which case I proceed to compute the roots numerically starting the search from a sufficiently low or from a sufficiently large starting point. If $V_j > U_i$ for this value of η , then $U(\bar{w}_i) \le V(P_j, Y_j^0)$.

2. Establish the feasibility conditions for each segment and kink as described in Eq. (9) in the paper, using θ_{jk} calculated in 1(a).

3. Combine the feasibility and optimality conditions for each segment and kink, as described in Section 5.1 in the paper, by taking the maximum of the lower bounds and the minimum of the upper bounds. Denote these values (L_k, H_k) and (l_k, h_k) for segments and kinks, respectively.

4. Substitute in the likelihood function as described in Appendix A.3.

5. Choose δ , α , γ , σ_{η} , and σ_{ε} to minimize the objective function. Iterate.

S5.2. Consistency

Let $W = (W_1, ..., W_N) \in \Theta$ denote the sequence of observations on the data and let G_N be the empirical distribution of W. Let L(b, W) denote the loglikelihood function derived in Appendix A.3, where $b = (\delta, \alpha, \gamma, \sigma_{\eta}, \sigma_{\varepsilon}) \in B$ denotes the vector of parameters to be estimated.

Manski (1988, Theorem 2') showed that the maximum likelihood estimator of b is consistent provided the following conditions hold:

CONDITION 1: There is a unique $b \in B$ for which $b = \arg \max_{c \in B} \int_{\Theta} L(c, W) dG$.

CONDITION 2: $L(\cdot, W)$ is continuous on B for all $W \in \Theta$.

CONDITION 3: There exists an integrable function $D: \Theta \to [0, \infty)$ for which $|L(c, W)| \le D(W)$ for all $(c, W) \in B \times \Theta$.

CONDITION 4: *B* is compact.

CONDITION 5: The observations W_i , $i = 1, ..., \infty$ are independent realizations from G.

Condition 1 states that the parameters $(\delta, \sigma_{\varepsilon}, \sigma_{\eta})$ are identified. This requires the assumption that the model (hence the conditional densities in the likelihood function) is well-specified. If this holds, as explained in the text, σ_{ε} and σ_{η} are identified from variation in consumption within and across segments (or kinks). Note that if the distribution of η is well-specified, it contains household-level heterogeneity without incurring an incidental parameters problem. In particular, household-level fixed effects are not estimated.

Condition 2 states that the likelihood function is continuous. In fact, as Manski (1988, Chapter 7.3) showed, this assumption can be relaxed to requiring that points of discontinuity have zero probability. Inspection of (20) shows that continuity of the likelihood function is satisfied.

Condition 3 is the condition for the Lebesgue dominated convergence theorem. In the standard case when the likelihood function is differentiable, this theorem guarantees that integration and differentiation of the likelihood function can be interchanged, which is used to show that the expected score is zero under the true parameters (e.g., Greene (2002, p. 475)). An important class of problems for which this condition fails is when the support of the dependent variable depends on the parameters. In my case, this might appear to be a problem because the support of η depends on the parameters due to the assumed truncation. However, note that the dependent variable also contains the optimization error ε which is distributed normally. Thus, the support of Wis $(-\infty, +\infty)$ regardless of the parameters.

Finally, Conditions 4 and 5 are conditions on the parameter space and the sample which we assume to hold. Note that this theorem does not require the differentiability of the likelihood function.

THE VALUE OF FREE WATER

S5.3. Uniqueness of the Demanded Quantity for any η

As long as preferences are convex, we know that demand exists and is unique for any kinked budget. (More precisely, uniqueness is true "almost surely," ignoring the case when a convex indifference curve has two tangency points with a non-convex part of the budget. Not to deal with this situation is standard.) This means that demand is unique for any η , and it is either w_1^0 , \bar{w}_1 , w_2^0 , or w_3^0 , since these are all the possibilities under the specific three-part budget considered here. The conditions given in (S1) are necessary for each of these cases to obtain:

$$(S1) \qquad w = \begin{cases} w_1^0 & \text{if } \begin{cases} (i) \ w_1^0 < \bar{w}_1 \text{ and} \\ (ii) \ V(P_1, Y) \ge V(P_3, Y_3^0) & \text{if } w_3^0 > \bar{w}_2, \\ \bar{w}_1 & \text{if } \begin{cases} (i) \ w_2^0 < \bar{w}_1 < w_1^0 \text{ and} \\ (ii) \ U(\bar{w}_1) \ge V(P_3, Y_3^0) & \text{if } w_3^0 > \bar{w}_2, \end{cases} \\ w_2^0 & \text{if } \begin{cases} (i) \ w_2^0 \in [\bar{w}_1, \bar{w}_2] \text{ and} \\ (ii) \ V(P_2, Y_2^0) \ge V(P_3, Y_3^0) & \text{if } w_3^0 > \bar{w}_2, \end{cases} \\ w_3^0 & \text{if } \begin{cases} (i) \ w_3^0 > \bar{w}_2 \text{ and} \\ (ii) \ V(P_1, Y) < V(P_3, Y_3^0) & \text{if } w_1^0 < \bar{w}_1 \text{ and} \\ (ii') \ V(P_2, Y_2^0) < V(P_3, Y_3^0) & \text{if } w_2^0 \in [\bar{w}_1, \bar{w}_2] \text{ and} \\ (ii'') \ V(P_2, Y_2^0) < V(P_3, Y_3^0) & \text{if } w_2^0 < \bar{w}_1 < w_1^0. \end{cases} \end{cases}$$

Each condition has two parts: (i) feasibility (consumption on the budget), and (ii) optimality (higher utility than the three other possibilities if they are feasible). In most cases, (ii) can be simplified, as done in (S1). For example, w_1^0 is demanded iff it is feasible and yields higher utility than \bar{w}_1, w_2^0 , and w_3^0 . In this case, feasibility of w_1^0 implies that neither \bar{w}_1 nor w_2^0 is feasible, so optimality simplifies to $V(P_1, Y) \ge V(P_3, Y_3^0)$. Clearly, these conditions are mutually exclusive (because of the optimality conditions). Therefore they are also sufficient for each case to obtain.

I now illustrate this with the specific functional forms resulting in Eq. (10) in the paper. That is, I show that, for any η , (10) uniquely defines a demanded quantity (without gaps or overlaps). Under convexity, we know that

$$\theta_{11} < \theta_{12} < \theta_{22},$$

where the second inequality follows from $\bar{w}_1 < \bar{w}_2$, and the first from the fact that $w_2^0 - w_1^0 = (P_2 - P_1)(\alpha + \gamma \bar{w}_1) < 0$ since $P_2 > P_1$ and $\alpha + \gamma \bar{w}_1 < 0$ from convexity. We do not know anything about θ_{23} , since w_3^0 can be anywhere on the extended budget constraint with P_3 . Thus, there are four possible

scenarios:

$$egin{aligned} & heta_{23} < heta_{11} < heta_{12} < heta_{22}, \ & heta_{11} < heta_{23} < heta_{12} < heta_{22}, \ & heta_{11} < heta_{23} < heta_{12} < heta_{22}, \ & heta_{11} < heta_{23} < heta_{12} < heta_{22}, \ & heta_{11} < heta_{12} < heta_{22} < heta_{23}. \end{aligned}$$

Consider the first one, and check that, for each possible η , (10) gives exactly one solution (and we can do the same for the other three scenarios). Denote the conditions in (10) for demand to be on the first segment, the kink, or the two other segments S1, K1, S2, and S3.

When $\theta_{23} < \theta_{11} < \theta_{12} < \theta_{22}$, we have the following possibilities (summarized in Figure S11):

If $\eta < \theta_{23}$: Only S1 holds (and since $\theta_{23} < \eta$ does not hold, the value of η_{13} is irrelevant). Observed consumption is predicted to be $w_1^0 + \eta + \varepsilon$.

If $\theta_{23} < \eta < \theta_{11}$: First part of S1 and S3 holds. If $\eta < \eta_{13}$, observed consumption is $w_1^0 + \eta + \varepsilon$. If $\eta > \eta_{13}$, observed consumption is $w_3^0 + \eta + \varepsilon$. Note that the second and third "and" in S3 are irrelevant since $\eta < \theta_{11}$.

If $\theta_{11} < \eta < \theta_{12}$: First part of K1 and S3 holds. Observed consumption is $\bar{w}_1 + \varepsilon$ if $\eta \in (u_{13}^L, u_{13}^U)$ and $w_3^0 + \eta + \varepsilon$ otherwise. The first and third "and" in S3 are irrelevant.

If $\theta_{12} < \eta < \theta_{22}$: First part of S2 and S3 holds. Observed consumption is $w_2^0 + \eta + \varepsilon$ if $\eta < \eta_{23}$ and $w_3^0 + \eta + \varepsilon$ otherwise. The first and second "and" in S3 are irrelevant.

If $\eta > \theta_{22}$: Only the first part of S3 holds, that is, $\theta_{23} < \eta$. The other parts of S3 are irrelevant. Observed consumption is $w_3^0 + \eta + \varepsilon$.

Figure S11 summarizes the consumer's observed consumption as a function of η for this scenario. The other three scenarios can easily be checked in the same way.

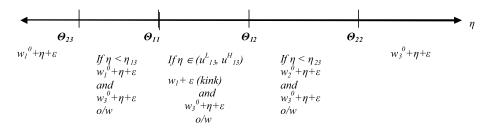


FIGURE S11.—Uniqueness of the demanded quantity for any η .

TABLE S-VII

EXPENDITURE CATEGORIES OF ODI WATER, THOUSAND RAND^a

| Bulk Water Purchase | 71,653 |
|----------------------|---------|
| Employee Cost | 64,375 |
| Operational Expenses | 152,337 |

^aAll data are from the 2010/2011 financial report of Odi Water (Odi Water Services (2011)). The document contains three major expenditure categories.

S6. COMPUTING THE MARGINAL COST OF WATER

All data are from the 2010/2011 financial report of Odi Water (Odi Water Services (2011)). The document contains three major expenditure categories (p. 12), which are shown in Table S-VII.

Employee Cost and Operational Expenses are unlikely to be directly connected to how much water is supplied to the households. It is difficult to judge how many new employees, if any, would need to be hired to provide more water to the same households; similarly, operational expenses include items like maintenance of the infrastructure, operating treatment facilities, etc., all of which are more closely related to the number of connections than to the volume of water supplied to these connections. In line with this interpretation, in the previous financial year, Odi Water purchased about 20% more bulk water, but had lower operational and employee costs.

Based on this, I compute the marginal cost of water in two alternative ways: first, I base it only on bulk water purchases; second, as a more "conservative" measure, I also include half of the operational expenses.

1. Marginal cost based on bulk water only

Odi Water receives almost all of its bulk water from Rand Water (a national bulk water supplier). Rand Water charges uniform water prices and Odi Water's water demand from Rand is negligible at a national scale so it is unlikely to affect this price. Based on the report (p. 16), Odi Water purchased 18,122,933 kiloliters of water in this fiscal year. I divide the cost of bulk water with this quantity to get the marginal cost of water, which is 4.0 Rand.

2. Marginal cost based on bulk water and operational costs

I take the marginal cost based on bulk water and add half of the operational costs divided by the amount of water purchased. This yields a marginal cost of 8.2 Rand.

S7. ESTIMATION RESULTS

S7.1. Marginal Effects and Price Elasticities

I average household-level marginal effects separately for various groups (indigent/non-indigent, restricted/non-restricted) to measure any differences

in the effects of explanatory variables across these groups. All marginal effects are based on the same set of household-level marginal effects estimated from the entire sample. The results are in Table S-VIII. The interpretation of the marginal effects of variables used to create the groups is as follows. For the indigent group, the marginal effect of the Indigent variable measures the effect of being indigent relative to the counterfactual of not being indigent. For the non-indigent group, it measures the effect of the counterfactual of becoming indigent relative to not being indigent. Table S-IX shows price elasticities by various consumer groups.

| | | | Groups | | |
|----------------------------------|----------|----------|--------------|------------|----------------|
| Variable | All | Indigent | Non-Indigent | Restricted | Non-Restricted |
| Price | -1.124 | -1.035 | -1.136 | -1.136 | -1.121 |
| | (0.003) | (0.013) | (0.002) | (0.007) | (0.003) |
| Income $\times 10^{-4}$ | 0.366 | 0.341 | 0.370 | 0.374 | 0.365 |
| | (0.004) | (0.015) | (0.005) | (0.011) | (0.004) |
| Avg. max daily temperature (°F) | 0.205 | 0.187 | 0.207 | 0.202 | 0.205 |
| | (0.001) | (0.003) | (0.001) | (0.002) | (0.001) |
| Number of people on the property | 0.055 | 0.051 | 0.056 | 0.055 | 0.055 |
| | (0.0005) | (0.002) | (0.0005) | (0.001) | (0.001) |
| Outdoor water usage | 0.102 | 0.092 | 0.103 | 0.100 | 0.102 |
| | (0.001) | (0.002) | (0.001) | (0.001) | (0.001) |
| Binary variables | | | | | |
| Indigent | 0.377 | 0.340 | 0.381 | 0.371 | 0.378 |
| - | (0.001) | (0.004) | (0.001) | (0.003) | (0.001) |
| Restricted | 0.357 | 0.330 | 0.360 | 0.350 | 0.358 |
| | (0.001) | (0.005) | (0.001) | (0.003) | (0.001) |
| Sanitation | 3.428 | 3.154 | 3.464 | 3.502 | 3.411 |
| | (0.027) | (0.074) | (0.032) | (0.065) | (0.037) |
| Washing machine | 0.093 | 0.085 | 0.094 | 0.094 | 0.093 |
| | (0.001) | (0.002) | (0.001) | (0.001) | (0.001) |
| Bathtub or shower | 5.814 | 5.655 | 5.834 | 5.929 | 5.787 |
| | (0.013) | (0.058) | (0.012) | (0.031) | (0.016) |
| Completed high school | -0.125 | -0.115 | -0.127 | -0.124 | -0.126 |
| | (0.001) | (0.002) | (0.001) | (0.002) | (0.001) |
| N | 10,000 | 1142 | 8858 | 1882 | 8118 |

TABLE S-VIII MARGINAL EFFECTS^a

^aFor continuous variables, the marginal effect reflects the impact of a unit increase in the variable on expected consumption (in kl). For categorical variables, it is the impact of an increase by one category (e.g., 0 to 1). For price, all marginal prices in the schedule are increased by 1 Rand. Expected consumption before and after the change is computed at the individual level as described in the Appendix of the paper, and averaged within the different consumer groups in each column. Standard errors are based on 100 bootstrapped samples of the same size as the estimation sample, taken with replacement. Reported standard errors are the standard deviations across these bootstraps.

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| | | Con | sumption Quartil | e | |
|---------------------------------|------------------------------|----------------------------|------------------------------|--------------------|--------------------------------|
| | 1st (1–6 kl) | 2nd (7–10 kl) | 3rd (11–17 kl) | 4th (18+ kl) | Overall |
| All | -1.022 | -0.997 | -0.971 | -0.923 | -0.976 |
| | (0.010) | (0.009) | (0.010) | (0.008) | (0.004) |
| Indigent | -0.954 | -0.964 | -0.879 | -0.873 | -0.915 |
| | (0.042) | (0.038) | (0.025) | (0.036) | (0.015) |
| Non-indigent | -1.030 | -1.001 | -0.984 | -0.930 | -0.984 |
| | (0.010) | (0.009) | (0.011) | (0.009) | (0.005) |
| Restricted | -1.049 | -1.016 | -0.983 | -0.935 | -0.935 |
| | (0.031) | (0.029) | (0.021) | (0.020) | (0.020) |
| Non-restricted | -1.017 | -0.993 | -0.968 | -0.920 | -0.973 |
| | (0.009) | (0.008) | (0.010) | (0.011) | (0.004) |
| Sanitation | -1.084 (0.012) | -1.021 (0.009) | -0.986 (0.010) | -0.935 (0.010) | -1.000 (0.005) |
| No sanitation | -0.806 (0.008) | -0.789 (0.011) | -0.762 (0.016) | -0.791 (0.014) | -0.793 (0.006) |
| Bathtub | -0.838 (0.013) | -0.849 (0.009) | -0.872 (0.010) | -0.839 (0.011) | -0.851 (0.006) |
| No bathtub | -1.137 (0.011) | -1.186 (0.013) | -1.186 (0.016) | -1.172 (0.018) | -1.166 (0.008) |
| Less than 2 persons on property | -1.095 | -1.060 | -1.002 | -0.907 | -1.035 |
| | (0.023) | (0.023) | (0.027) | (0.024) | (0.013) |
| More than 2 persons on property | -1.005 | -0.985 | -0.967 | -0.925 | -0.967 |
| | (0.011) | (0.010) | (0.009) | (0.009) | (0.005) |
| Below average income | -1.037 (0.010) | -1.030 (0.012) | -0.984 (0.010) | -0.957 (0.013) | -1.004 (0.006) |
| Above average income | -0.983 | -0.938 | -0.954 | -0.886 | -0.933 |
| | (0.019) | (0.015) | (0.015) | (0.013) | (0.007) |
| Completed high school | -1.0208 | -0.9922 | -0.9598 | -0.9129 | -0.968 |
| | (0.012) | (0.010) | (0.011) | (0.011) | (0.005) |
| No completed high school | -1.0259 (0.018) | -1.007 (0.015) | -0.9984 (0.018) | -0.9513 (0.015) | -0.996 (0.007) |
| Washing machine | -0.995 | -0.957 | -0.948 | -0.911 | -0.947 |
| | (0.011) | (0.010) | (0.010) | (0.010) | (0.005) |
| No washing machine | (0.012) -1.050 (0.013) | (0.014) -1.054 (0.014) | (0.010) -1.010 (0.015) | -0.951 (0.015) | (0.000) (-1.020) (0.008) |

TABLE S-IX PRICE ELASTICITIES BY CONSUMER GROUPS^a

^aPrice elasticities reflect the percentage change in consumption in response to a 1 percent change in all marginal prices (0 prices are increased to 1 Rand). Elasticities are based on expected consumption before and after the change, computed at the household level as described in the Appendix of the paper, and averaged within the different consumer groups and consumption quartiles. Standard errors are based on 100 bootstrapped samples of the same size as the estimation sample, taken with replacement. Reported standard errors are the standard deviations across these bootstraps.

S7.2. Model Performance

Table S-X investigates the out-of-sample performance of the model. I use the estimated parameters to predict consumption for the 53,178 monthly observations that were not used in the estimation.

| MODEL PERFORMANCE: OUT OF SAMPLE ^a | | | | | | | |
|---|-------------|----------------|---------------|--------|--|--|--|
| | Actual Mean | Predicted Mean | Average Error | N | | | |
| All | 13.395 | 13.063 | -0.332 | 53,178 | | | |
| Indigent | 13.206 | 13.369 | 0.163 | 6392 | | | |
| Non-indigent | 13.421 | 13.021 | -0.400 | 46,786 | | | |
| Restricted | 14.990 | 13.687 | -1.303 | 10,349 | | | |
| Non-restricted | 13.010 | 12.912 | -0.098 | 42,829 | | | |

^aEntries in the 'Actual' and 'Predicted' columns are average household consumption levels in kl. The Predicted mean column gives the average expected consumption predicted by the model with the estimated parameter values in the Appendix of the paper. Expected consumption is computed at the individual level as described in the Appendix of the paper. Average error is the difference between the actual and predicted means. Values in the table are for the 53,178 monthly observations that were not used in estimating the parameters.

TABLE S-X MODEL PERFORMANCE: OUT OF SAMPLE

S8. COUNTERFACTUAL ANALYSIS WITHOUT FREE WATER

TABLE S-XI

COUNTERFACTUAL ANALYSIS WITHOUT FREE WATER: HOUSEHOLD CONSUMPTION AND EXPENDITURE CHANGES, ALTERNATIVE EFFECTIVE PRICES^a

| | A 11 | T | Non Indianat | Destricted | N Dt-i-t-d |
|-----------------------------|---------|----------|--------------|------------|----------------|
| | All | Indigent | Non-Indigent | Restricted | Non-Restricted |
| Consumption (kl/month) | | | | | |
| With free water | 13.307 | 13.408 | 13.293 | 13.826 | 13.183 |
| Without free water, | | | | | |
| 30% higher effective prices | 12.391 | 12.220 | 12.414 | 12.814 | 12.291 |
| Change (%) | -6.880 | -8.166 | -6.708 | -3.702 | -7.632 |
| Without free water, | | | | | |
| 60% higher effective prices | 9.939 | 9.318 | 10.021 | 10.242 | 9.868 |
| Change (%) | -25.307 | -29.974 | -24.691 | -23.030 | -25.841 |
| Without free water, | | | | | |
| 90% higher effective prices | 3.849 | 2.778 | 3.989 | 2.901 | 4.075 |
| Change (%) | -71.074 | -79.123 | -70.022 | -78.199 | -69.376 |
| Expenditures (Rand/month) | | | | | |
| With free water | 77.337 | 80.140 | 76.971 | 85.033 | 75.506 |
| Without free water, | | | | | |
| 30% higher effective prices | 147.693 | 152.527 | 147.062 | 159.801 | 144.812 |
| Change (%) | 90.973 | 97.224 | 90.157 | 106.629 | 87.248 |
| Without free water, | | | | | |
| 60% higher effective prices | 131.264 | 129.694 | 131.468 | 142.271 | 128.644 |
| Change (%) | 69.730 | 67.700 | 69.994 | 83.962 | 66.342 |
| Without free water, | | | | | |
| 90% higher effective prices | 67.685 | 62.671 | 68.339 | 64.277 | 68.496 |
| Change (%) | -12.480 | -18.964 | -11.635 | -16.887 | -11.432 |

^aValues reported are the model predicted values using either the actual water tariffs ("With free water"), or the counterfactual tariffs where 0 prices were replaced with 30, 60, and 90 percent higher prices than the actual effective price ("Without free water"). Expected consumption is computed at the individual level as described in the Appendix of the paper, and averaged within the different consumer groups in each column. Expenditure is average household water spending (in 2008 Rand). N = 7309 (observations after the 2007 policy change are excluded from both the actual and the counterfactual computations).

S9. OPTIMAL TARIFFS

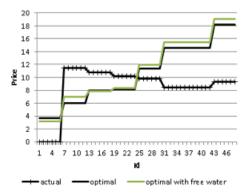


FIGURE S12.—Optimal and actual 2006/2007 tariff schedules, low marginal cost. *Notes*: The figure shows the actual 2006/2007 tariff and the corresponding optimal tariffs with and without 6 kl free water for indigent households. The marginal cost is set to 4 Rand. The "optimal with free water" schedule shown is for non-indigent households.

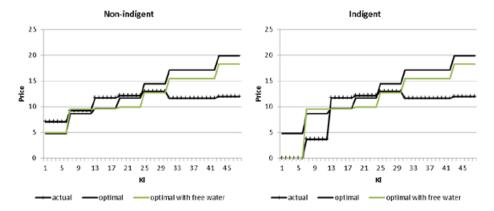


FIGURE S13.—Optimal and actual 2008/2009 tariff schedules, low marginal cost. *Notes*: The figure shows the actual 2008/2009 tariff and the corresponding optimal tariffs with and without 6 kl free water for indigent households. The marginal cost is set to 4 Rand.

| | 2006/2007 Op | timal Tariff | | 2008/2009 Op | timal Tariff | |
|-----------------------------|--------------------|----------------|------|--------------------|-----------------|------|
| | Without Free Water | Wih Free Water | Ν | Without Free Water | With Free Water | Ν |
| Panel A: High marginal cost | | | | | | |
| Compensating variation | | | | | | |
| All | -3.98/-6.59 | -3.61/-5.00 | 8385 | -1.96/-12.54 | -1.04/-7.51 | 5660 |
| Indigent | -3.75/-6.27 | -25.66/-28.48 | 1021 | 61.59/61.79 | 37.35/37.61 | 877 |
| Non-indigent | -4.01/-6.62 | -0.55/-3.43 | 7364 | -13.62/-12.79 | -8.09/-8.07 | 4783 |
| Consumption | | | | | | |
| Consumption (kl) | 15.87/16.57 | 16.04/17.02 | 8385 | 13.36/14.11 | 12.85/13.87 | 8385 |
| Low consumption (%) | 19.11 | 15.56 | - | 41.27 | 41.48 | - |
| Medium consumption (%) | 49.48 | 49.83 | _ | 58.66 | 58.52 | - |
| High consumption (%) | 31.41 | 34.61 | - | 0.07 | 0.00 | - |
| Panel B: Low marginal cost | | | | | | |
| Compensating variation | | | | | | |
| All | -18.51/-21.78 | -18.03/-20.51 | 8385 | -8.80/-19.20 | -7.88/-13.15 | 5660 |
| Indigent | -18.35/-21.29 | -34.70/-37.92 | 1021 | 54.87/55.55 | 30.53/31.05 | 877 |
| Non-indigent | -18.54/-21.29 | -15.72/ -19.29 | 7364 | -20.48/-20.22 | -14.92/-14.18 | 4783 |
| Consumption | | | | | | |
| Consumption (kl) | 16.81/17.63 | 16.75/17.74 | 8385 | 14.19/14.93 | 13.90/14.98 | 8385 |
| Low consumption (%) | 11.87 | 11.56 | - | 38.60 | 38.22 | - |
| Medium consumption (%) | 44.00 | 45.27 | - | 61.13 | 56.57 | - |
| High consumption (%) | 44.13 | 43.17 | - | 0.27 | 5.21 | - |
| Panel C: Actual consumption | | | | | | |
| Consumption (kl) | 13.36/1 | 14.49 | 8385 | 12.76/1 | 2.60 | 8385 |
| Low consumption (%) | 55.0 | 50 | - | 65.4 | 14 | - |
| Medium consumption (%) | 21.1 | 13 | _ | 17.9 | 91 | - |
| High consumption (%) | 23.2 | 27 | _ | 16.7 | 75 | - |

 TABLE S-XII

 COMPENSATING VARIATION UNDER THE OPTIMAL TARIFFS (MEAN/MEDIAN)^a

^aThe table reports the compensating variation corresponding to the optimal tariffs. If the provider switched from the actual (2006/2007 or 2008/2009) tariff to the optimal tariff, this is the change in income that would leave a consumer as well off as he was before the switch. Negative numbers indicate an increase in consumer utility from the switch. In each cell, the first entry is the mean, the second entry is the median compensating variation. All entries are in 2008 Rand. Each row in the table presents the fraction of consumers consuming Low (0-12 kl), Medium (12-18 kl), or High (above 18 kl) quantities of water under the tariff schedule indicated in the first column.

| | Consum | ption (kl) | Consumption Distribution (Percent) | | | |
|---------------------------------------|--------|------------|------------------------------------|--------|-------|-------|
| Tariff Schedule | Mean | Median | Low | Medium | High | Total |
| Panel A: actual | | | | | | |
| 2006/2007 | 13.36 | 14.49 | 55.60 | 21.13 | 23.27 | 100 |
| 2008/2009 | 12.76 | 12.6 | 65.44 | 17.81 | 16.75 | 100 |
| Panel B: optimal (high marginal cost) | | | | | | |
| 2006/2007 optimal | 15.87 | 16.57 | 19.11 | 49.48 | 31.41 | 100 |
| 2006/2007 optimal with free water | 16.04 | 17.02 | 15.56 | 49.83 | 34.61 | 100 |
| 2008/2009 optimal | 13.36 | 14.11 | 41.27 | 58.66 | 0.07 | 100 |
| 2008/2009 optimal with free water | 12.85 | 13.87 | 41.48 | 58.52 | 0.00 | 100 |
| Panel C: optimal (low marginal cost) | | | | | | |
| 2006/2007 optimal | 16.81 | 17.63 | 11.87 | 44.00 | 44.13 | 100 |
| 2006/2007 optimal with free water | 16.75 | 17.74 | 11.56 | 45.27 | 43.17 | 100 |
| 2008/2009 optimal | 14.19 | 14.93 | 38.60 | 61.13 | 0.27 | 100 |
| 2008/2009 optimal with free water | 13.90 | 14.98 | 38.22 | 56.57 | 5.21 | 100 |

TABLE S-XIII

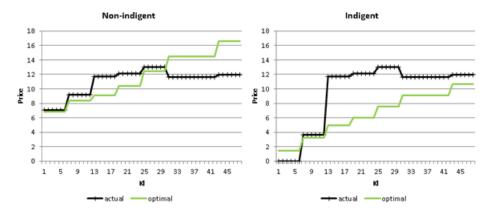
DISTRIBUTION OF CONSUMPTION UNDER ACTUAL AND OPTIMAL TARIFF SCHEDULES^a

^aEach row in the table presents the fraction of consumers consuming Low (0–12 kl), Medium (12–18 kl), or High (above 18 kl) quantities of water under the tariff schedule indicated in the first column.

| | 2006/2007 | 5/2007 2006/2007 Optimal Tariff | | 2008/2009 | 2008/2009 Op | timal Tariff | | |
|-----------------|-------------|---------------------------------|-------------------|-----------|--------------|-----------------------|--------------------|------|
| | Actual | Without Free Water | Wih Free Water | Ν | Actual | Without Free Water | With free Water | Ν |
| Panel A: High m | arginal co | st | | | | | | |
| All | 86.13 | 103.06 | 105.06 | 8385 | 93.34 | 101.43 | 97.16 | 5660 |
| Indigent | 81.85 | 102.49 | 82.37 | 1021 | 48.45 | 100.31 | 70.67 | 877 |
| Non-indigent | 86.76 | 103.14 | 108.37 | 7364 | 101.57 | 101.63 | 102.02 | 4783 |
| Restricted | 101.41 | 106.13 | 106.47 | 2011 | 108.94 | 109.24 | 103.06 | 1208 |
| Non-restricted | 81.35 | 102.09 | 104.62 | 6374 | 89.11 | 99.31 | 95.56 | 4452 |
| Panel B: Low m | arginal cos | st | | | | | | |
| All | 86.13 | 96.54 | 96.53 | 8385 | 93.34 | 102.73 | 100.95 | 5660 |
| Indigent | 81.85 | 96.01 | 79.33 | 1021 | 48.45 | 101.39 | 74.39 | 877 |
| Non-indigent | 86.76 | 96.62 | 98.92 | 7364 | 101.57 | 102.98 | 105.83 | 4783 |
| Restricted | 101.41 | 99.51 | 98.12 | 2011 | 108.94 | 109.34 | 105.94 | 1208 |
| Non-restricted | 81.35 | 95.61 | 96.03 | 6374 | 89.11 | 100.94 | 99.60 | 4452 |

TABLE S-XIV Expenditure Under the Optimal Tariffs^a

 a The table reports average expenditures under the optimal tariffs discussed in Section 7.3. It also reports the observed average expenditure under the actual 2006/2007 and 2008/2009 tariffs for comparison. All entries are in 2008 Rand.



S10. SEPARATE TARIFFS FOR INDIGENT AND NON-INDIGENT HOUSEHOLDS

FIGURE S14.—Separate optimal and actual 2008/2009 tariff schedules for non-indigent and indigent households, low marginal cost. *Notes*: The figure shows the optimal and actual 2008/2009 tariff when the social planner sets the tariff for the non-indigent group only (left-hand side) or the indigent group only (right-hand side). Optimal tariffs maximize welfare subject to a profit neutrality constraint for the given group only. The marginal cost is set to 4 Rand.

TABLE S-XV

Compensating Variation and Consumption Under Indigent-Only and Non-Indigent-Only Optimal ${\rm Tariffs}^{\rm a}$

| | Compensa | Compensating Variation | | Consumption (kl) | | |
|-------------------|------------|------------------------|-------|------------------|----------|--|
| | Mean | Median | Mean | Median | Ν | |
| Panel A: High mar | ginal cost | | | | <u> </u> | |
| Indigent | -2.66 | -7.14 | 16.63 | 18.00 | 877 | |
| Non-indigent | -2.74 | -2.00 | 13.72 | 14.75 | 4783 | |
| Panel B: Low marg | ginal cost | | | | | |
| Indigent | -17.64 | -22.39 | 18.42 | 19.04 | 877 | |
| Non-indigent | -11.79 | -11.91 | 14.65 | 15.60 | 4783 | |

^aThe table reports the compensating variation corresponding to the indigent-only and non-indigent-only optimal tariffs, as well as the consumption resulting from these tariffs. If the provider switched from the actual (2008/2009) tariff to the optimal tariff for the given group, the compensating variation is the change in income that would leave a consumer in this group as well off as he was before the switch. Negative numbers indicate an increase in consumer utility from the switch.

TABLE S-XVI

EXPENDITURE UNDER INDIGENT-ONLY AND NON-INDIGENT-ONLY OPTIMAL TARIFFS^a

| | 2008/2009 Actual Tariff | Optimal Tariff | N |
|--------------------------|-------------------------|----------------|------|
| Panel A: High marginal c | cost | | |
| Indigent | 48.45 | 64.61 | 877 |
| Non-indigent | 101.57 | 116.05 | 4783 |
| Panel B: Low marginal c | ost | | |
| Indigent | 48.45 | 61.33 | 877 |
| Non-indigent | 101.57 | 115.59 | 4783 |

^aThe table reports average expenditures under the optimal tariffs discussed in Section 7.5. It also reports the observed average expenditure under the actual 2008/2009 tariffs for comparison. All entries are in 2008 Rand.

S11. OPTIMAL TARIFFS UNDER REVENUE AND CAPACITY CONSTRAINTS

TABLE S-XVII

COMPENSATING VARIATION UNDER THE OPTIMAL TARIFFS WITH CAPACITY AND REVENUE NEUTRALITY CONSTRAINTS (MEAN/MEDIAN)^a

| | Optimal | Tariff | |
|----------------|--------------------|-----------------|------|
| | Without Free Water | With Free Water | Ν |
| All | -0.40/-11.88 | -0.94/-7.41 | 5660 |
| Indigent | 63.52/63.14 | 37.69/38.01 | 877 |
| Non-indigent | -12.05/-12.04 | -8.03/-8.03 | 4783 |
| Restricted | 7.07/-11.80 | 3.60/-7.14 | 1208 |
| Non-restricted | -2.43/-11.90 | -2.18/-7.50 | 4452 |

^aThe table reports the compensating variation corresponding to the optimal tariffs discussed in Section 7.6 of the paper. If the provider switched from the actual tariff to the optimal tariff, this is the change in income that would leave a consumer as well off as he was before the switch. Negative numbers indicate an increase in consumer utility from the switch. In each cell, the first entry is the mean, the second entry is the median compensating variation. All entries are in 2008 Rand.

TABLE S-XVIII

CONSUMPTION DISTRIBUTION UNDER THE OPTIMAL TARIFFS WITH CAPACITY AND REVENUE NEUTRALITY CONSTRAINTS^a

| | Consumption (kl) | | Con | Consumption Distribution (Percent) | | | |
|-------------------------|------------------|--------|-------|------------------------------------|-------|-------|--|
| Tariff Schedule | Mean | Median | Low | Medium | High | Total | |
| Actual | 12.76 | 12.6 | 65.44 | 17.81 | 16.75 | 100 | |
| Optimal | 12.76 | 13.00 | 41.41 | 58.59 | 0 | 100 | |
| Optimal with free water | 12.76 | 13.74 | 41.41 | 58.59 | 0 | 100 | |

^aThe table describes consumption under the actual and optimal tariffs. For the distribution of consumption, Low, Medium, and High refer to the fraction of consumers consuming, respectively, 0–12 kl, 12–18 kl, or above 18 kl.

TABLE S-XIX

| | 2008/2009 | Optimal | Optimal Tariff | | |
|----------------|---------------|--------------------|----------------|------|--|
| | Actual Tariff | Without Free Water | Wih Free Water | Ν | |
| All | 93.34 | 96.38 | 96.32 | 5660 | |
| Indigent | 48.45 | 95.99 | 70.15 | 877 | |
| Non-indigent | 101.57 | 96.45 | 101.12 | 4783 | |
| Restricted | 108.94 | 103.57 | 102.24 | 1208 | |
| Non-restricted | 89.11 | 94.42 | 94.72 | 4452 | |

EXPENDITURE UNDER THE OPTIMAL TARIFFS WITH CAPACITY AND REVENUE NEUTRALITY CONSTRAINTS^a

^aThe table reports average expenditures under the optimal tariffs discussed in Section 7.6. It also reports the observed average expenditure under the actual 2008/2009 tariffs for comparison. All entries are in 2008 Rand.

S12. ROBUSTNESS TO ALTERNATIVE INCOME MEASURES

To investigate the robustness of the model estimates to the income measure used, I also estimated the model using only sampled households with both reported and estimated income. Specifically, from my random 10,000 estimation sample, I pulled households with observed income data (N = 5731). Next, I estimated the model for these 5731 household-month observations separately, using either estimated or observed income. The resulting parameter vectors from the likelihood model are in Table S-XXI. Next, I computed the predicted consumption and average model error for these two cases separately, and I also simulated the average price elasticity in both cases. These are included in Table S-XX, and the numbers are consistent with those used in the main text. This is not too surprising since the estimated coefficient on the income variable is small in each case, and household income (even reported income) has a relatively small variance in this sample.

| | Observed | Estimated | Results From the |
|----------------------------|----------|-----------|------------------|
| | Income | Income | Main Text |
| Actual mean consumption | 13.276 | 13.276 | 13.353 |
| Predicted mean consumption | 13.033 | 13.012 | 13.071 |
| Average price elasticity | -0.243 | -0.264 | -0.282 |
| | -0.979 | -0.981 | -0.976 |
| N | 5731 | 5731 | 10,000 |

TABLE S-XX

MODEL PERFORMANCE AND PRICE ELASTICITY ESTIMATES WITH ALTERNATIVE INCOME MEASURES^a

^aThe table reports model performance measures and estimated average price elasticities using alternative samples and income measures. The first two columns show ML estimates using only the 576 households whose reported income is non-missing in the survey. The first column uses reported income, while the second column uses estimated income for these households in the estimation. The last column shows the corresponding values from the main text.

| | Observed | d Income | Estimate | d Income | Results From | the Main Text |
|---------------------|-----------------------|-----------------------|-----------------------|-----------------------|-----------------------|------------------------|
| Variables | Parameter | SE | Parameter | SE | Parameter | SE |
| Constant | 2.134 | 0.028 | 2.115 | 0.026 | 2.111 | 0.011 |
| Indigent | 0.357 | 0.175 | 0.366 | 1.917 | 0.364 | 0.142 |
| Restricted | 0.341 | 0.024 | 0.343 | 0.071 | 0.344 | 0.033 |
| Sanitation | 4.793 | 0.046 | 4.789 | 0.100 | 4.787 | 0.038 |
| Average max daily | | | | | | |
| temperature (°F) | 0.197 | 0.000 | 0.197 | 0.001 | 0.197 | 0.000 |
| Washing machine | 0.090 | 0.039 | 0.091 | 0.054 | 0.091 | 0.014 |
| Number of people | | | | | | |
| on the property | 0.054 | 0.007 | 0.053 | 0.021 | 0.053 | 0.003 |
| Outdoor water | | | | | | |
| usage | 0.099 | 0.009 | 0.098 | 0.053 | 0.097 | 0.007 |
| Bathtub or shower | 6.255 | 0.031 | 6.317 | 0.034 | 6.261 | 0.013 |
| Completed high | | | | | | |
| school | -0.117 | 0.021 | -0.121 | 0.079 | -0.120 | 0.044 |
| Price | -1.140 | 0.001 | -1.139 | 0.014 | -1.139 | 0.002 |
| Income | $0.366 	imes 10^{-4}$ | $0.060 	imes 10^{-4}$ | $0.347 	imes 10^{-4}$ | $0.057 	imes 10^{-4}$ | $0.358 	imes 10^{-4}$ | 0.046×10^{-4} |
| σ_η | 0.005 | 0.017 | 0.005 | 0.006 | 0.005 | 0.006 |
| σ_{ϵ} | 9.167 | 0.220 | 9.202 | 0.220 | 9.233 | 0.163 |
| N | 5731 | | 5731 | | 10,000 | |

TABLE S-XXI PARAMETER ESTIMATES WITH ALTERNATIVE INCOME MEASURES^a

^aThe table presents parameter estimates from the maximum likelihood model using alternative samples and income measures. The first two sets of parameters were estimated using only the 576 households whose reported income is non-missing in the survey. The first two columns use reported income, while the next two columns use estimated income for these households in the estimation. The last set of parameter estimates are those reported in the main text.

S13. ROBUSTNESS TO CONTROLLING FOR FIRST TARIFF QUARTER

Some of the graphs in Section S4 suggested that consumers' response to tariff changes may be different in the first tariff quarter. In this section, I investigate whether controlling for this affects the estimates reported in the paper. Specifically, I repeat the estimation including in the vector Z a dummy variable equal to 1 for observations in the first quarter of each tariff year (July–September) and 0 otherwise. This allows for any differences in consumer behavior in this quarter.

The resulting parameter estimates are in Table S-XXII. For comparison, the table also reports the estimates from the main text. As can be seen, the coefficient on the newly included First Quarter dummy is not statistically significant. To investigate the performance of the model, Table S-XXIII compares actual and predicted mean consumption, again listing the corresponding values from the main text for comparison. As can be seen, although the average performance of the model improved somewhat, it is now less accurate in predicting consumption *both* in the first quarter and in other tariff quarters. Average performance only improved because the negative average error in predicting first

| | | First Dummy | Results From the Main Text | | |
|-----------------------|-----------------------|------------------------|-------------------------------|------------------------|--|
| Variables | Parameter | SE | Parameter | SE | |
| Constant | 2.999 | 0.089 | 2.111 | 0.011 | |
| Indigent | 0.269 | 0.073 | 0.364 | 0.142 | |
| Restricted | 0.194 | 0.089 | 0.344 | 0.033 | |
| Sanitation | 5.250 | 0.061 | 4.787 | 0.038 | |
| Average max daily | | | | | |
| temperature (°F) | 0.198 | 0.002 | 0.197 | 0.000 | |
| Washing machine | 0.069 | 0.051 | 0.091 | 0.014 | |
| Number of people | | | | | |
| on the property | 0.042 | 0.006 | 0.053 | 0.003 | |
| Outdoor water usage | 0.112 | 0.024 | 0.097 | 0.007 | |
| Bathtub or shower | 6.512 | 0.057 | 6.261 | 0.013 | |
| Completed high school | -0.138 | 0.101 | -0.120 | 0.044 | |
| First quarter | -0.190 | 0.131 | _ | _ | |
| Price | -1.259 | 0.024 | -1.139 | 0.002 | |
| Income | $0.308 	imes 10^{-4}$ | 0.043×10^{-4} | $0.358 	imes 10^{-4}$ | 0.046×10^{-4} | |
| σ_η | 0.005 | 0.008 | 0.005 | 0.006 | |
| σ_{ϵ} | 9.281 | 0.161 | 9.233 | 0.163 | |
| Ν | 10,000 | | 10,000 | | |

TABLE S-XXII

PARAMETER ESTIMATES CONTROLLING FOR FIRST TARIFF QUARTER^a

^aThe table presents parameter estimates from the maximum likelihood model including an indicator for observations from the first tariff quarter (July–September). For comparison, the table also shows the parameter estimates reported in the main text.

TABLE S-XXIII

| | With First Quarter Dummy | Results From the Main Text |
|----------------------------|-----------------------------|-------------------------------|
| All quarters | | |
| Actual mean consumption | 13.353 | 13.353 |
| Predicted mean consumption | 13.201 | 13.071 |
| Average error | -0.152 | -0.282 |
| N | 10,000 | 10,000 |
| Quarter 1 | | |
| Actual mean consumption | 13.180 | 13.180 |
| Predicted mean consumption | 12.266 | 12.311 |
| Average error | -0.914 | -0.869 |
| N | 2471 | 2471 |
| Quarters 2–4 | | |
| Actual mean consumption | 13.410 | 13.410 |
| Predicted mean consumption | 13.508 | 13.321 |
| Average error | 0.098 | -0.089 |
| N | 7529 | 7529 |

MODEL PERFORMANCE CONTROLLING FOR FIRST TARIFF QUARTER^a

^aEntries in the 'Actual' and 'Predicted' rows are average household consumption levels in kl. Predicted mean is the average expected consumption predicted by the model with the estimated parameter values. Expected consumption is computed at the individual level as described in the Appendix of the paper. Average error is the difference between the actual and predicted means. The table presents these values for all tariff quarters as well as separately for the first and other tariff quarters. The first column uses the new parameter estimates including the first quarter dummy; the second column uses the parameter estimates from the paper for comparison.

quarter consumption is now partially offset by a positive average error in predicting consumption in other quarters.

Overall, these results do not suggest that differences in consumer behavior in the first tariff quarter are statistically significant, nor that controlling for them necessarily improves the fit of the model.

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Dept. of Economics, University of Houston, 3623 Cullen Boulevard, Houston, TX 77204, U.S.A.; aszabo2@uh.edu.

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