Social Discounting and Externalizing Behavior Problems in Boys

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ABSTRACT

Social discounting refers to the fact that most people assign more value to the welfare of close affiliates than they do to the welfare of distant affiliates—they *discount* the latter compared to the former. We report the first study to apply a social discounting paradigm to boys. We were particularly interested in investigating the relations between social discounting, age, and externalizing behavior problems (antisocial behavior). Results showed that (1) preadolescent boys were more likely than adolescent boys to show atypical response patterns in allocating rewards to affiliates; (2) task behavior was well represented as social discounting once boys with atypical response patterns were deleted from the sample, and (3) boys functioning in the clinical range on indices of externalizing behavior problems demonstrated steeper social discounting compared to controls. We conclude that social discounting as a measure of perceived social closeness is feasible for use in adolescent samples. Social discounting may operate similarly to other forms of discounting in impulsive individuals. Copyright © 2011 John Wiley & Sons, Ltd.

KEY WORDS social discounting; externalizing behavior problems; adolescents

INTRODUCTION

Elevated impulsivity underlies many externalizing disorders in children and adolescents (Quay, 1993). Externalizing disorders refer to a group of disorders characterized by antisocial behavior, aggression, rule breaking, impulsivity, and overactivity (Achenbach & Edelbrock, 1978; Hill, 2002; Hinshaw, 1992). Externalizing behavior problems may be distinguished from internalizing problems characterized by depression, withdrawal, dysphoria, and anxiety (Quay, 1986). A leading framework for understanding impulsivity is temporal discounting, which refers to the tendency of people to attach less value to rewards that are distant in time than to rewards that are close in time. Widely cited literature in psychology and behavioral economics maintains that such discounting is common in people for many or most reward classes (Ainslie, 1992, 2001; Herrnstein & Prelec, 1992). In this literature, the default intertemporal discount function for a particular individual is given by a hyperbolic function as described by Mazur's (1987) formula (1):

$$v = \frac{A}{1+kD} \tag{1}$$

where v, A, and D represent, respectively, the present value of a delayed reward, the amount of a delayed reward, and the delay of the reward. The parameter k is a constant that may differ among individuals; the greater is k, the greater the discounting by delay on reward value for that individual. In contrast to the exponential discounting functions traditionally used by economists to model attitudes toward delayed monetary rewards, hyperbolic intertemporal discounting allows for intertemporal preference reversals when agents choose between smaller, sooner rewards (SSRs) and larger, later ones (LLRs). For instance, say a choice involves an SSR of \$6 and an LLR of \$10. If the \$6 SSR is available immediately and the \$10 LLR is available in one week, many people would prefer the \$6 SSR. However, if the delay to both rewards is increased by one year, so the \$6 SSR is available in one year and the \$10 LLR is available in one year plus one week, virtually everyone would now prefer the \$10 LLR. Hyperbolic discounting much more readily accounts for such preference reversals than does exponential discounting. The construct of intertemporal discounting is regularly used to model impulsive behavior, for example, struggles to maintain diets, opportunistic and regretted marital infidelities, and consumption of addictive substances (Ainslie, 1975, 1992, 2001).

Jones and Rachlin (2006) applied hyperbolic discounting to model people's valuations of social affiliations at a single point in time. They hypothesized that people's ranking of acts of generosity toward specific others in their social networks is well described by a function modeled after (1) above. This can be represented as

$$v_{i,j} = \frac{A_i}{1 + k' N_{i,j}} \tag{2}$$

where $v_{i,j}$, represents the value person *i* attaches to the welfare of a particular affiliated person *j*; A_i , represents the value she associates with her own welfare; and $N_{i,j}$ is the rank person *i* assigns to a given affiliate *j* among *i*'s full list of affiliated people. The constant k' may differ among individuals; the greater is k', the greater the discounting by social distance on reward value for that individual.

In Jones's and Rachlin's experiment, participants were instructed to think of an imaginary list ranked in order of social closeness as numbers 1, 2, 5, 10, 20, 50, and 100.

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Therefore, the 1st person on the list would be the person closest to them (e.g., husband or wife) and the 100th person would be a more distant acquaintance (e.g., cashier at the grocery store). They were then asked whether they would prefer to receive a certain amount of money (e.g., \$155) for themselves only (the "selfish" option) or \$75 for themselves and \$75 for a designated person on their list (the "share" option). The amounts for selfish options descended from \$155 to \$75 in decrements of \$10 (thus each subject imagined nine choices). They were asked to imagine these choices for persons in order of the affiliates they had ranked as numbers 1, 2, 5, 10, 20, 50, and 100 on the list.

Jones and Rachlin found that, unsurprisingly, participants were willing to forgo greater amounts of money for the benefit of people with whom they perceived themselves as having closer affiliation. For example, some subjects were willing to forgo the entire \$75 for a person high on their list, but only willing to forgo \$10 for a person low on their list. Jones and Rachlin (2006) assumed that subjects compared the rewards they received when they chose selfishly with the rewards they gave to others on internally consistent linear scales. The authors then showed that under that assumption, if all choices by all subjects are aggregated, the data best fit a hyperbolic curve (Equation 2). This was reconfirmed in a second study with different subjects using a similar experimental design (Rachlin & Jones, 2008a). Rachlin and Jones (2008a, b) therefore interpreted the choice behavior displayed by their experimental subjects as "hyperbolic social discounting."

A motivation of Rachlin and Jones (2008b) in exploring this phenomenon is their general hypothesis that people comparatively value objects, relationships, and experiences in a way that implies hyperbolic distance measures from reference points. Thus, according to Jones and Rachlin, intertemporal discounting and social discounting-along with discounting of risk-are special cases of a general human disposition to associate greater relative ratios of value to comparisons among higher-valued alternatives than to comparisons among lower-valued alternatives. This would allow for comparisons of relative dispositions to selfish choice between subjects by reference to differences in the values of their personal k' parameters. That is, a subject with a higher k' parameter will reduce his generosity as social distance increases at a faster rate than a subject with a lower *k'* parameter.

A question raised by consideration of this hypothesis is whether there is evidence that it applies to children, who are stereotypically viewed as impulsive by comparison with normal adults. Limited investigations of intertemporal discounting in children have been reported. While most researchers expect discounting to be represented by a hyperbolic discount curve that is comparable in functional form to those of adults, children are also expected to discount delayed rewards more steeply. Indeed, Green, Fry, and Myerson (1994) acquired experimental data showing that Mazur's formula yielded a close fit to discounting in 12-year-old, while Scheres et al. (2006) showed that younger children (6–11 years old) discounted delayed rewards more steeply than did older children (12–17 years old). Green et al. (1994) also demonstrated steeper k-values for children.

Against this background, the main aim of the present study was to investigate whether the social discounting paradigm developed by Jones and Rachlin (2006) could be applied to children and adolescents (aged 8–17) taking into account externalizing behavior problems. First, we determined whether boys' choices for different affiliate rankings match intuitive expectations about conceptualization of social space on which the Jones and Rachlin paradigm is based. Do boys choose individuals who are indeed close to them in social distance, or do they make choices that suggest difficulty in understanding the abstract idea of social distance? Imagine, for example, that boys consistently assign high ranks to people they see regularly, but with whom they have few or no reciprocal personal interactions, for example, the janitor at school.

Second, we investigated whether any unexpected (atypical) patterns in social discounting in boys might be accounted for by age. As observed by Jones and Rachlin, a typical adult pattern of social discounting is choosing the "share" option for closest affiliates, but switching to the "selfish" option at some point down the list as more distant affiliates are considered. Jones's and Rachlin's experimental paradigm, however, incorporates features which may be sensitive to developmental effects. It requires participants to hold in mind seven relationships with specific others, by reference to which nine decisions each are made (see Method Section for more detail). It is possible that younger children lack adequate executive functioning (Pennington & Ozonoff, 1996), working memory or abstract thinking capacity to reliably complete the task, and might therefore make random and inconsistent decisions between options.

Third, we investigated the extent to which implementation of the Jones and Rachlin paradigm with child and adolescent subjects produces patterns of implied social discounting that resemble Jones's and Rachlin's adult data in respect of being well described by Mazur's formula. Positive findings in this regard would further establish the applicability of the social discounting paradigm to youth.

Fourth, we investigated whether steeper social discounting (that is, social discounting best characterized by functions in which the k' parameter is higher valued) is related to indices of externalizing behavior disorders in boys. Children and adolescents with externalizing problems have a greater preference for small and immediate rewards compared to larger delayed rewards (Barkley, Edwards, Laneri, Fletcher, & Metevia, 2001; Kuntsi, Oosterlaan, & Stevenson, 2001; Schweitzer & Sulzer-Azaroff, 1995; Solanto et al., 2001; Sonuga-Barke, Taylor, Sembi, & Smith, 1992; Tripp & Alsop, 2001; for a review, see Luman, Oosterlaan, & Sergeant, 2005), indicating greater impulsivity in these children. Children and adolescents with externalizing behavior problems have also been shown to have significant social-cognitive deficits (Dodge, 1980, 1993; Sharp, 2008) and problems with empathy (Blair, 1997, 1999). Higher impulsivity and problems in social cognition and empathy would suggest that children functioning in the clinical range on indices of externalizing behavior problems might make choices best described by social discounting functions with systematically higher k' values, compared to children below cut-off.

METHOD

Participants

A sample of 170 boys (2nd to 12th graders) was recruited through community organizations in Houston, Texas. We focused the study on boys because of the known high risk of externalizing behavior problems associated with male gender (Hill, 2002; Moffitt, Caspi, Rutter, & Silva, 2001).

The mean age and grade of the sample was 162 months (13.5 years) and 7th grade, respectively, with the youngest aged 8 (n = 1) and the oldest aged 17 (n = 10). Just under half of the boys (48.8%; n = 83) were between ages 8–12. The sample was ethnically representative with socio-economic status estimated to be primarily middle class on the basis of parental education. Parents and children consented in person prior to data collection. IRB approval for the study was obtained.

Measures

Externalizing behavior problems

It is well known that different sources (e.g. parents, peer, and children themselves) often disagree on the presence and severity of problem behaviors (Verhulst & van der Ende, 1992), each contributing valid and unique information (Achenbach, McConaughy, & Howell, 1987; Verhulst & Van der Ende, 1992). Multiple informants are needed to obtain a comprehensive picture of an individual's functioning (Verhulst & van der Ende, 1992). To this end, we combined three measures of externalizing behavior to include youth self-report, parent-report, and peer nominations.

Youth self-report and parent-report. The youth self-report (YSR; Achenbach & Rescorla, 2001) and child behavior checklist (CBCL; Achenbach & Rescorla, 2001) are well established evidence-based assessment instruments (Holmbeck et al., 2008), that assess global and more specific psychopathology among boys ages 6-18 years. The YSR and CBCL assess behavioral and emotional disorders in the past 6 months. The YSR and CBCL each contains 112 problem items, each scored on a 3-point scale 0 (not true), 1 (somewhat or sometimes true), to 2 (very or often true). The YSR and CBCL each yields a T-score of general psychiatric functioning and two broad subscales of externalizing behavior problems and internalizing behavior problems. Achenbach and Rescorla (2001) recommended using a T-score cut-off at or above 65 to separate individuals at higher risk for psychopathology. Prior research indicates this threshold discriminates well between clinical and nonclinical populations (Achenbach & Rescorla, 2001). For the current study, we used T-score of 65 on the externalizing subscale to group boys in above- and below-cut off categories for clinically relevant externalizing behavior problems.

Peer nomination as relationally aggressive or prosocial. A peer-nomination instrument developed by Werner and Crick (1999) was used to assess relational aggression and prosocial behavior. The measure consists of 24 items. Seven of the items tap into a relational aggression subscale. This subscale has been found to be highly reliable, with Cronbach's alpha of 0.87. Nine items tap into a prosocial behavior subscale that has been shown to be highly reliable as well (Cronbach's alpha = 0.91).

Following Werner and Crick (1999), participants were provided with a group membership roster to be used during self-administration of the peer-nomination instrument. For each of the behavioral items, participants were instructed to nominate up to five peers who best fit each description. The number of nominations each participant received from his or her peers was summed for each item and totaled for each subscale.

Next, we used the 50th percentile to identify boys above and below the median for relationally aggressive and prosocial nominations. As is often the case with sociometric studies (Cillessen & Bukowski, 2000) many boys were nominated as both prosocial and relationally aggressive, so many boys fell above the 50th percentile for both subscales. For the analyses, we were interested in comparing social discounting in boys who were perceived as "only prosocial" versus boys nominated in both categories. Similarly, we were interested in comparing boys who were "only relationally aggressive" to boys in both categories.

Composite measure of externalizing behavior problems. Since normality assumptions for externalizing behavior disorder variables were violated (possibly due to the fact that the sample was drawn from the community), all variables were submitted to a normalizing transformation. A boy was considered to meet criteria for externalizing behavior problems if he were above clinical cut-off for parent- and self-report externalizing problems, considered as "only relationally aggressive," and not in the "only prosocial category."

Social discounting

The social discounting measure used by Jones and Rachlin (2006) was adapted for use in the current study. The top page of the boys' questionnaire packet informed participants that their answers were anonymous and would remain confidential, that there were no right or wrong answers, and that honest answers were preferred. The instructions were as follows: The following game asks you to *imagine* that you have made a list of the 100 people closest to you in the world ranging from your dearest friend or relative at position #1 to someone you barely know at #100.

The person at number one would be someone you know well and is your closest friend or relative.

The person at #100 might be someone you recognize and encounter but perhaps you may not even know their name. You do not have to physically create the list—*just imagine that you have done so*. On the next few pages you will be asked to make a series of decisions based on what you prefer. Your choice will always be between

A. keeping money for yourself or

B. sharing money with someone else

Participants were also asked to raise their hands if they needed help. On the first page of the instrument, participants were again reminded of their task:

Imagine you made a list of the 100 people closest to you in the world ranging from your dearest friend or relative at #1 to someone you barely know at #100.

On each of the seven pages of the measure appeared the following instructions. (We show as an example the instruction applied with respect to the first person on each subject's list.)

Write down the name of the first person on your list.

What is the relationship you have to that person (e.g., mom, dad, sister, brother, best friend, boyfriend, girlfriend, etc.)?

Now imagine the following choices between an amount of money for you and an amount for the first person (#1) on the list.

Circle A or B to indicate which you would choose in *EACH* line.

A. \$155 for you alone.	B. \$75 for you and 75 for the first
	person on the list.
A. \$145 for you alone.	B. \$75 for you and 75 for the first
	person on the list.
A. \$135 for you alone.	B. \$75 for you and 75 for the first
	person on the list.
A. \$125 for you alone.	B. \$75 for you and 75 for the first
	person on the list.
A. \$115 for you alone.	B. \$75 for you and 75 for the first
	person on the list.
A. \$105 for you alone.	B. \$75 for you and 75 for the first
	person on the list.
A. \$95 for you alone.	B. \$75 for you and 75 for the first
	person on the list.
A. \$85 for you alone.	B. \$75 for you and 75 for the first
	person on the list.
A. \$75 for you alone.	B. \$75 for you and 75 for the first
	person on the list.

For all participants, the left column (column A) was identical on all pages. For each participant, the social distance ("first" above) of the right column (column B) was the same on a single page, but varied between pages, changing from 1 to 100, in the following increments and order: 1, 2, 5, 10, 20, 50, and 100. That way, participants were indicating how much money they were willing to forgo for themselves alone so they and a person on their list could both receive money.

Procedures

The social discounting and peer nomination tasks were presented to groups of 5–10 children with three trained experimenters in the room. The small group administration

enabled participants to ask clarification questions about the instructions. Trained experimenters took care not to suggest preferred ways of responding to the task. Parents completed the CBCL during the same testing session but in a different room.

Data analytic strategy

For data analysis, a single crossover point monetary value where participants switched from choosing the option in the A column (the selfish option) to choosing the option in the B column (the share option) was determined for each subject following Jones and Rachlin (2006). The difference between the crossover point value $x: 575 \le x \le 155$ (where x varies with social distance point) and \$75 was the maximum amount of money a participant would forgo in exchange for \$75 to the person at each social distance point on the list. The crossover point was defined as the mean of column A dollar amounts where the switch between columns A and B was made. For example, if a participant chose the "selfish" option at A: \$155 and A: \$145 and switched to the "share" option at A: \$135, the crossover point was calculated to be \$140. However, many participants chose the share option even when the alternative was A: \$155. In these cases, following Jones and Rachlin (2006), a crossover point was assumed at \$160. When a participant chose the selfish option exclusively for a given position on her list, a crossover point was assumed at \$70, again after Jones and Rachlin (2006).

To arrive at a discount rate for different individuals, formula (2) was fitted to each subject's crossover point data using a least-squares criterion. In general, that is, the sum of squares $\sum_{n=1}^{7} (v_{i,n} - \hat{v}_{i,n})^2$ was minimized with respect to both A_i and k' for each case, where $v_{i,n}$ is the crossover point at social distance $N_{i,n}$. This process generated values of A_i and k'_i for each child. In this particular analysis we assumed that the A_i were constant (at A_0) across individuals and also assumed that the seven social distances for each individual were the same but that k'_i differs from person to person as usual. We thus estimated a function of the following form, using a least squares criterion as described above:

$$v_{i,n} = \frac{A_0}{1 + k'_i N_n}$$
(3)

In order to estimate A_0 , we proceeded as follows: we first found the median indifference point for each of the seven social distances; these seven medians were then used to fit the Mazur hyperbolic discounting function to find A and $k'_{average}$; the constant A obtained in this (median) case was labeled A_0 . We then estimated Equation (3) for each individual *i* and assuming a constant A (A_0) we obtained a k'_i for each individual using a least squares fitting procedure.

RESULTS

Children's understanding of social distance, as reflected in choices of people at varying positions

An important way of determining whether the social discounting paradigm could usefully be administered to children was to determine whether subjects' choices for



Figure 1. Percentage of subjects' (n = 170) choices of persons to occupy each of the first four social distance points (1st, 2nd, 5th, and 10th)

affiliate rankings were intuitively consistent with the expectations underlying Jones's and Rachlin's design. As shown in Figure 1, subjects' choices of persons to occupy each of the first four social distance points mirrored conventional expectations. For example, 36.3% of children designated a parent and 36.8% designated a best friend as first person on the list (i.e., closest in social distance), compared to siblings (9.4%), girlfriends (6.4%), other family members (2.9%) and others (1.8%). In contrast, only 11.7% of children designated a parent as second on the list, with 57.3% designating a friend at this position and 9.4, 6.4, 8.2, and 2.9% of children choosing the other categories of designated persons displayed in Figure 1. The patterns of choices for the fifth and tenth positions on the list remained similar to those of the second position.

The relationship between response patterns and age

More than half of the sample (66.5%; n = 113) made inconsistent responses with multiple crossover points (randomly switching between "share" decisions and "selfish" decisions regardless of affiliates' rankings) or "perverse" responses—that is, crossover points in the wrong direction (choosing the "selfish" option for higher-ranked affiliates and "share" options for lower-ranked affiliates). We tested for a relationship between atypical decisionmaking of the kinds described immediately above, and age (n = 170). The results of a chi-square analysis showed that 74% of 8–12 versus 40% of 13–17 years old showed atypical response patterns, thereby demonstrating a significant likelihood for younger children to engage in atypical response patterns, $\chi^2(1, N = 170) = 3.59$; p = 0.05).

Are children's hypothetical sharing choices with affiliates well described as social discounting?

Before plotting subjects' behavior using formula (2), it was necessary to exclude participants with atypical or inconsistent responses. Criteria for exclusion included multiple crossover points and crossover points in the wrong direction, as described above. Table 1 summarizes the percentage Table 1. Percentage inclusion by age group (exclusion based on multiple switch-over points and switch-over points in the wrong direction; n = 170)

Exclu			
Yes	No	Total	
1 100	0 0	1 100	
4 100	0 0	4 100	
33 82.5	7 17.5	40 100	
23 60.5	15 39.5	38 100	
21 70.0	9 30.0	30 100	
13 68.4	6 31.6	19 100	
9 45.0	11 55.0	20 100	
4 50.0	4 50.0	8 100	
5 50.0	5 50.0	10 100	
113 66.5	57 33.5	170 100	
		Excluded Yes No $1 \\ 100$ $0 \\ 0$ $4 \\ 100$ $0 \\ 0$ $33 \\ 100$ $7 \\ 17.5$ $23 \\ 60.5$ 17.5 $23 \\ 60.5$ 39.5 $21 \\ 70.0$ $9 \\ 30.0$ $13 \\ 68.4$ $6 \\ 68.4$ $9 \\ 45.0$ 51.0 $4 \\ 50.0$ 50.0 $5 \\ 50.0$ 50.0 $5 \\ 50.0$ 50.0 $113 \\ 66.5$ $57 \\ 33.5$	

Table 2. Social distances, k' and R^2 (least squares fit) in the case where A is fixed (at A_0 on subjects who were included in the study (n = 57; exclusion based on multiple switch-over points and switch-over points in the wrong direction)

Subject	Social distances							Estimates (hyperbolic fit)		
	1.0	2.0	5.0	10.0	20.0	50.0	100.0	A_0	k'	$R^{2}(\%)$
1	75	75	75	75	65	55	5	93.72	0.029	70.53
2	75	75	75	75	45	45	5	93.72	0.042	79.83
3	25	35	15	15	15	15	5	93.72	1.384	-35.57
4	85	75	65	45	35	25	15	93.72	0.088	96.74
5	55	45	85	35	15	5	5	93.72	0.179	52.09
6	85	85	85	85	35	5	5	93.72	0.055	85.22
7	85	85	75	65	55	65	55	93.72	0.014	-26.75
8	85	85	65	45	35	45	5	93.72	0.073	84.22
9	55	45	85	85	5	5	5	93.72	0.101	40.00
10	85	85	85	85	45	5	5	93.72	0.05	86.00
11	25	5	15	5	5	5	5	93.72	3.021	42.53
12	85	85	55	15	5	5	5	93.72	0.195	89.35
13	75	75	55	45	25	15	15	93.72	0.126	96.09
14	85	85	85	75	23 75	45	5	93.72	0.026	87.04
15	15	15	15	5	5	5	5	93.72	3 487	-19.62
16	55	35	25	25	5	5	5	03.72	0.644	89.65
17	95 95	95 95	25	25	95 95	95 95	65	02 72	0.044	22.05
17	6J 55	6J 45	0J 25	0J 25	0J 25	0J 25	5	93.72	0.004	25.75
18	33 75	45	33 25	25 25	25	25	5	93.72	0.397	45.51
19	15	15	35	33	25	5	5	93.72	0.202	94.48
20	85	45	25	55	35	25	25	93.72	0.182	14.41
21	85	22	35	85	85	25	15	93.72	0.037	22.41
22	5	5	15	15	5	5	5	93.72	8.328	214.13
23	85	75	75	75	25	5	5	93.72	0.076	87.62
24	75	75	75	75	25	25	25	93.72	0.059	77.12
25	85	85	85	85	15	5	5	93.72	0.067	80.09
26	85	85	75	85	35	25	5	93.72	0.05	88.64
27	85	85	85	85	45	5	5	93.72	0.05	86.00
28	5	85	85	75	5	5	5	93.72	0.106	15.35
29	85	85	65	45	45	25	45	93.72	0.059	59.32
30	25	15	45	55	5	5	5	93.72	1.135	-48.65
31	25	15	5	5	5	5	5	93.72	2.676	84.88
32	65	65	45	25	25	15	5	93.72	0.231	92.20
33	85	85	85	25	5	5	5	93.72	0.132	83.29
34	85	75	85	5	45	45	35	93.72	0.08	25.23
35	85	85	85	65	55	45	45	93.72	0.023	73.48
36	25	25	25	25	15	5	5	93.72	1.443	-65.01
37	75	65	65	65	55	35	5	93.72	0.053	72.98
38	75	75	75	75	75	45	5	93.72	0.029	73 71
30	85	85	85	85	85	35	15	93.72	0.022	84 50
40	85	85	85	85	85	25	5	93.72	0.022	80.51
40	85	85	85	85	85	85	55	03.72	0.027	55 72
42	85	85	85	85	65	55	25	93.72	0.005	92.87
42	85	85	65	5	5	5	25 5	03.72	0.019	92.07
43	85	85	85	85	5	15	5	93.72	0.195	76.46
44	0J 05	0J 05	0J 05	0J 05	25	15	5	93.72	0.009	70.40
45	83 85	83 85	83 95	6J 05	55	23	5	95.72	0.047	00.13 70.07
40	85	85	85	85	5	35	5	93.72	0.059	12.21
4/	5	5	85	85	5	5	5	93.72	1.064	-43.60
48	85	85	85	85	55	45	25	93.72	0.025	93.68
49	85	85	65	35	35	25	25	93.72	0.09	87.48
50	85	85	75	75	5	25	5	93.72	0.077	80.39
51	85	85	85	85	85	35	25	93.72	0.02	85.66
52	25	25	5	5	5	5	5	93.72	2.145	76.92
53	55	45	25	15	5	5	5	93.72	0.592	98.26
54	85	85	85	85	85	55	45	93.72	0.011	88.02
55	5	55	65	5	5	5	5	93.72	0.819	-9.01
56	55	65	35	35	15	5	5	93.72	0.295	85.75
57	85	85	85	85	45	5	5	93.72	0.05	86.00
Average								93.72	0.536	55.29



Figure 2. Data plot with hyperbolic and exponential fit of median crossover points. The median crossover points for each social distance was fitted to a hyperbolic curve using the least squares criterion. Best fit was shown to be with the hyperbolic curve

exclusions per age group. It is clear that younger children were over-represented in the exclusion category compared with older children.

The final sample size for this analysis was n = 57, which accounted for 33% of the original sample. The mean age of the final sample size was M = 13.54 (SD = 1.85), with the youngest child aged 11 (n = 7) and the oldest aged 17 (n = 5).

The results of fitting formula (2) to each subject's crossover point data using a least-squares criterion are listed in Table 2. The average k' was 0.115, and the average fit R^2 was 74.9%.

It will be seen from Table 2 that several cases constitute outliers and distort the results, often by inflating k'. Thus, for example, subject #31 gives very low crossover points and a very high resultant k'. To more robustly estimate the k' across the group, we took the median crossover points for each social distance and fitted a hyperbolic curve using the least squares criterion as before to these (median) points (see Figure 2). This fit gave a k' of 0.79.

For comparability with Jones and Rachlin (2006) and Rachlin and Jones (2008b), we show the data fit to both hyperbolic and exponential curves. The function for the latter is

$$v_{i,j} = A_i e^{-k_i N_{i,j}} \tag{4}$$

where e is the base of the natural logarithm. Figure 2 clearly shows that the fit to the hyperbolic curve was superior.

Social discounting and externalizing behavior problems in children

The *k*-values of children above cut-off on the composite externalizing behavior problems variable were compared to those for children below cut-off. An independent sample *t*-test with the composite score (above and below cut-off) as independent variable and *k*-values as the dependent variable showed the variances for these two groups to be unequal (Levene's test for equality of variances, F = 4.57; p = 0.03). The alternative *t*-value which compensates for unequal

variance was significant, t(55) = 2.43; p = 0.01, indicating a moderate effect size ($\eta^2 = 0.09$). Boys in the clinical range for externalizing behavior problems therefore demonstrated significantly higher k'-values (steeper discounting rates) compared to boys functioning within normal range. To ensure that age was not confounding the difference between above- and below cut-off boys on social discounting, we ran an independent sample *t*-test. The mean age for boys above cut-off (m = 13.91; SD = 1.51) was not significantly different from the mean age of boys below cut-off (m = 13.46; SD = 1.92), t(55) = -0.72, p = 0.47.

DISCUSSION

This is the first study to investigate social discounting in nonadults (boys), and the first to investigate the relationship between social discounting and externalizing behavior problems. As such, this study extends previous work using the social discounting paradigm in adults (Jones & Rachlin, 2006, 2009; Osinski, 2009, 2010; Rachlin and Jones, 2008a, b).

Findings suggest that preadolescent children (ages 11 and under) engage in social discounting behavior that is difficult to rationalize, and therefore, the use of the Jones and Rachlin paradigm is premature with this age range. It is possible that the experimental task simply asks too much of younger children's executive functioning (Pennington & Ozonoff, 1996), working memory, or abstract thinking capacity. However, all children's choices of affiliates were intuitively appropriate, suggesting that their problem was not the ability to understand or operationalize the idea of ranking affiliations. Discriminating patterns of generosity against these rankings among adolescents (ages 12 and up) were well represented by a hyperbolic social discounting function. Social discounting showed a significant relationship with externalizing behavior problems in that children above cutoff on a variety of indices of externalizing behavior problems demonstrated steeper social discounting compared to children below cut-off.

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Limitations of the study suggest future work. A clinical sample would be predicted to show a stronger relationship between social discounting and psychopathology. A larger sample size, restricted to adolescents, would permit more powerful statistical analysis. Moreover, future work in younger children may adapt the procedure to be more developmentally sensitive, for instance, by developing a more straightforward task, possibly an oral version with fewer and easier questions, and lower monetary values involved.

We offer no comment on the ambitious Jones and Rachlin (2006) hypothesis that people manifest a general disposition to discount all values hyperbolically from reference points. It is independently interesting to observe that adolescents hypothetically allocate monetary rewards to others in a way consistent with hyperbolic discounting over "affiliation space." This suggests that the k' parameter in the social discounting function may usefully measure an element of, or correlate with, forms of psychopathology in children.

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